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CLIMATE CHANGE AND AGRICULTURAL DEVELOPMENT IN WEST AFRICA: ROLE OF RENEWABLE ENERGY AND TRADE OPENNESS

Abstract

The design, implementation, and evaluation of energy policies in combating climate change are becoming increasingly evident to strengthen economic growth driven by the agricultural sector in most developing countries. The study analyzes the direct and indirect effects of renewable energy consumption (REC) on agriculture value-added (AgVA), CO₂ emissions, and trade openness in the short- and long-run in the West African countries. The second-generation panel unit root tests, the panel cointegration methods, and Panel Vector Error Correction Model are used with World Bank data from 1990 to 2015. A panel Granger causality test was also used to determine the direction of causality between variables. Findings show a unidirectional relationship between AgVA, CO₂ emissions, and REC; between REC, gross fixe capital formation (GFCF) and trade openness. Moreover, the bidirectional hypothesis is verified between agricultural development and trade openness. However, the null hypothesis is found between AgVA and GFCF, on the one hand, and GFCF and CO₂ emissions, on the other hand. These results suggest that fostering renewable energy policy and revisiting trade policy toward reducing environmental pollution will enable agricultural development and boost the regional economy.

Keywords

renewable energy, agriculture, CO₂ emissions, trade openness, panel cointegration, West Africa

JEL Classification

O13, Q27, Q56

INTRODUCTION

The role of renewable energy (RE) in economic development and its environmental benefits in terms of climate risk management has increased interest in recent debates around the world (Bayale et al., 2021; Frangou et al., 2018; Rafindadi & Ozturk, 2017; Liu et al., 2017a). According to Liu et al. (2017a) and Heidari and Pearce (2016), RE can be a key instrument in climate change adaptation and mitigation. It is well recognized that CO₂ emissions using RE technology are less than traditional energy supply sources (Liu et al., 2017a; Ben Jebli & Ben Yousef, 2015; Heidari & Pearce, 2016). Increasing investment in RE production and consumption could be more economically beneficial and more viable than non-renewable energy use (Frangou et al., 2018; Kahia et al., 2017; Rafindadi & Ozturk, 2017). For example, Frangou et al. (2018) have estimated that saving from renewable energy consumption (REC) could be ranged from 3% to 23% on energy costs in the case study of Greece. According to Miketa and Merven (2013), the share of RE technologies will increase from 22% to 52% of current electricity generation in West Africa, with 46% of adding capacity by 2030. The use of RE as part of sustainable development goals is said to contribute significantly to poverty reduction and countries' develop-

ments. The REC could boost the rural households' economies and enhance job creation opportunities, while the non-access to energy would severely affect economic growth (Schwerhoff & Sy, 2017; Inglesi-Lotz, 2016; Birol, 2007).

Despite many studies in energy economics, there are still debates on the effect and relationship direction between REC and economic growth. Four hypotheses are often discussed:

- (I) the neutrality assumption, which does not support any relationship between energy use and economic growth;
- (II) the unidirectional assumption between economic growth and energy consumption;
- (III) the growth hypothesis, which strongly agrees that energy consumption leads to economic growth; and
- (IV) feedback assumption, which assumes that energy consumption positively affects economic growth, and vice versa (Brini et al., 2017; Ben Jebli & Ben Youssef, 2017).

The specific energy policy could depend not only on the type of hypothesis but also on the geographical position of the region and the main drive of countries' economies. In the context of globalization, trade openness that eases RE technology transfer can also lead to new empirical insights and policy implications. While the agricultural sector remains the main wagon of West African countries' economies (more than 35% of GDP), not much attention has been given to investigating the relationship between REC, CO₂ emissions, and agricultural development in the energy economics of the region. Moreover, while trade openness can be a catalyst in RE transfer, it can be a source of pollution in the context of globalization in regions with non-binding environmental laws such as West African countries, and therefore, more investigations are needed.

This study fills this research gap in the empirical literature by analyzing the dynamics between climate change captured by CO₂ emissions and agricultural development in West African countries while highlighting the role of renewable energy and trade openness. Specifically, this research assesses the short and long-run effects of REC and trade openness on agricultural development in West African countries. It also investigates the causality direction between REC, CO₂ emissions, and agricultural development in the short- and long-run in the study areas. The results will help foster the design and implementation of an energy policy that encourages the creation and promotion of small-scale enterprises that work to develop renewable energy technology in the region. This study is also in line with the Sustainable Development Goals (SDG) agenda that seeks to ensure efficiency in terms of affordability and accessibility of energy by 2030. The study will increase agricultural productivity in the region through energy use while combating climate change by reducing CO₂ emissions.

1. LITERATURE REVIEW

Natural resources such as energy must be managed so that future generations could benefit, while its consumption is expected to meet the needs of the population as defined sustainable development by Brundtland (1987). Indeed, the theory of sustainable development might find its origin after Meadows et al. (1972) demonstrate that economic growth could be limited if resources are not ra-

tionally or efficiently used. Since then, the theory of green growth and sustainable development in the production process has gained researchers' attention in the context of climate change and agricultural development (Zaccour & Oubraham, 2018; Reilly, 2012). It is societies' and policymakers' responsibility to think about the effects of today's actions on future generations. This refers to the concept of energy efficiency. The use of RE is then highly motivated for agricultural sustainability in

a green growth perspective (Chel & Kaushik, 2011; Adelaja & Hailu, 2008). It is an example of agricultural land and other factors such as environment, biomass, and water resources. For instance, agriculture, which is the main source of production and livelihood of most of the population, contributes significantly to the CO₂ emissions (Liu et al., 2017a; Ben Jebli & Ben Youssef, 2015). This leads to the rethinking of sustainability in the agricultural production process, hence the green economy concept. The green economy concept includes the trade-off between RE and non-renewable energy consumption that would have a strong link with economic growth (Bayale et al., 2021; Lyytimäki, 2018; Sutherland et al., 2015). Martinho (2018) and Sutherland et al. (2015) showed that agriculture has an important role in the green economy as a key sustainable development strategy. Energy security relies on the affordability of energy but must consider the technological change (Proskuryakova, 2018).

The economic advantage that offers REC and the volatile nature of energy prices affecting global economies led to countries' setting up the roadmaps of 100% of REC (Sadiqa et al., 2018; Hohmeyer & Bohm, 2015). REC must be encouraged in developing countries, given the rapid population growth and urbanization that could increase energy demand, which is forecasted to increase by 100% on the global scale and more than 120% in some countries, such as Russia and Brazil by 2050 (Resch et al., 2008). This is evident in West African countries where the population is expected to increase by 2.5% per annum by 2050, increasing CO₂ emissions from human activities and electricity generation technologies from non-renewable energy sources. Despite the contribution of energy to economic growth, the International Renewable Energy Agency (IRENA, 2018) has reported that the current state of energy consumption is alarming in West African countries. For instance, the average rate of electricity access per capita ranges between 9% and 72%, with most countries less than 30% (IRENA, 2018) compared to 98% in North Africa (IRENA, 2015). Also, the International Energy Agency (IEA) (2014) has reported that West Africa is the only region in the world where the number of people living without electricity is increasing. Investing in RE technology adoption in the region could help reduce the

energy gap and boost economic growth driven by the agriculture sector.

Adelaja and Hailu (2008) found that using wind energy in the Michigan agricultural sector could increase farm net revenue by 50%. Similarly, Paramati et al. (2018) found that REC has positively impacted economic activities, including agriculture, more than non-renewable energy. Whatever it comes from a renewable or non-renewable source, the energy constitutes a key input in the agricultural production process. During crop production and processing, RE use can directly or indirectly play an important role (transport of agricultural inputs, harvesting, packing, seeding and irrigation, poultry production, and transformation of animals' derived products). It can also help manage post-harvest losses (grain drying, food transformation, storage, and conservation). As a result, agriculture and its value chain actors could be highly vulnerable to the variability of fuel prices on the international market (Farajian et al., 2018; Martinho, 2016) if based production more on non-renewable energy consumption. Turkey's case study by Bayrakci and Koçar (2012) showed a high trade surplus estimated at EUR 1.5 billion because of REC in agriculture.

Renewable energy use is an encouraging option in combating climate change induced by an increase in CO₂ emissions from fossil fuel consumption. For example, agriculture is seen as the main source of CO₂ emissions. Simultaneously, using RE will reduce CO₂ emissions and increase agricultural value-added higher than the conventional energy impact (Wesseh & Lin, 2016). The countries' dependency on fossil fuel imports and the increase in power costs because of rising demand are other reasons that support RE policy redesign, especially in Sub-Saharan Africa, where 90% of the population still do not have access to electricity (IRENA, 2018). For instance, the use of fossil fuels, pesticides, and synthetic fertilizer in agriculture could increase production costs and destroy progressively agricultural ecosystems and increase environmental damage with decreasing farm profitability in the long run because of the degradation of natural soil nutrients. Chel and Kaushik (2011) found that the use of wind energy could reduce CO₂ emissions by approximately 0.9 tons per year. For sustainable agriculture, the

use of RE could combat pollution, encourage the use of clean technologies, and facilitate the use of water pumps for irrigation (Chel & Kaushik, 2011).

Whether theoretically or empirically, the energy policy and recommendations for economic development among decision-makers are not unanimous. The results depend on whether one considers the short- or long-run patterns, the area of study, or the methods adopted for the typical study. Using the Granger causality test, Ben Jebli and Ben Youssef (2017) found the unidirectional causality between economic growth and REC and agricultural value-added, non-REC and agricultural value-added, and between CO₂ emissions and RE in the short run. Moreover, the feedback hypothesis was found between agricultural value-added and CO₂ emissions, while the U-shaped environmental curve hypothesis was not verified. The bidirectional assumption was not supported in the case study of OECD countries (Alp, 2016). Using the vector autoregressive model, Johansen cointegration, and Granger causality test, Alp (2016) found heterogeneous results between countries within the OECD region. There was no evidence of the relationship between energy consumption and economic development in eleven countries, while the growth hypothesis that strongly agrees on the energy consumption and economic growth nexus was satisfied within 6 countries only (Alp, 2016).

The heterogeneity characteristic of the effects of RE on economic growth was also found by Bhattacharya et al. (2016). A significant and positive effect of RE on economic development was found for 57% of the selected countries (Bhattacharya et al., 2016). Also, using quarterly time-series data, Rafindadi and Ozturk (2017) point out that REC, capital, and labor productivity could strengthen the German economy. The bidirectional effects between RE and economic growth were found. Considering the multivariate panel framework, Kahia et al. (2017) found the long-run relationship between REC, non-renewable energy consumption, labor force, and gross capital formation over 1980–2012. Bidirectional causality was found between REC and economic growth, non-renewable energy consumption, and economic growth using the panel error correction model (Kahia et al., 2017). However, not much attention has been given to the nexus between climate change, REC,

trade openness, and agricultural development within the West African countries that promote regional integration for several decades. The study will expand the literature to foster regional energy policy for sustainable development.

2. METHODS

Several methods have been used to assess the socioeconomic impact of REC and environmental management. The methods used to analyze the energy demand and supply and the impact assessment of energy transition depend on whether one considers the national or regional scales (Farajian et al., 2018; Jenniches, 2018; Khan et al., 2018; Bhattacharya et al., 2016). The production function is often served as the theoretical foundation in analyzing REC and economic growth nexus (Bhattacharya et al., 2016; Inglesi-Lotz, 2016). The model applied depends on the econometric tests required and the data used. Based on the green growth theory developed by Hickel and Kallis (2020), the neoclassical production function in which some inputs are used to produce a certain level of output is used. Assume that the agricultural development captured by the agricultural value-added (*AgVA*) is mainly affected by greenhouse gas emissions (*CO2*), trade openness (*Trade*), and renewable energy consumption (*REC*) while controlling for other variables (*X*). Then, it follows:

$$AgVA_{it} = f(REC_{it}, CO2_{it}, Trade_{it}, X_{it}, \varepsilon_{it}). \quad (1)$$

Trade openness (*Trade*) represents the ratio of exports and imports to GDP. Gross fixed capital formation (GFCF) representing domestic investment is used as a control variable. In equation (1), *i* represents the country (*i* = 1... 12) and *t* is the year (*t* = 1990... 2015). The choice of variables is based on the literature and the study period depends on the availability of the data. Assuming that the specific functional form of equation (1) is the Cobb-Douglas production function, equation (1) could be rewritten in the linear form as follows:

$$\begin{aligned} \ln AgVA_{it} = & \vartheta_i + \varphi_{1i} \ln CO2_{it} + \\ & + \varphi_{2i} \ln REC_{it} + \varphi_{3i} \ln Trade_{it} + \\ & + \varphi_{4i} GFCF_{it} + \delta_i t + \varepsilon_{it}, \end{aligned} \quad (2)$$

where ε_{it} is the error term, \mathcal{G}_i is a column vector capturing the country-specific effect, and t is the deterministic time trend lasting from 1990 to 2015. φ and θ are the vectors of parameters to be estimated.

The cross-sectional dependence within individuals in panel data may be occurred because of common shocks due to a strong economic integration within countries (for example, the regional energy policy of ECOWAS, the common agricultural policy in WAEMU (ECOWAS and WAEMU are sub-regional blocs within the West African countries), coronavirus disease (COVID-19) or spatial dependence (WAEMU countries versus non-WAEMU countries; two sub-regional groups within ECOWAS). Ignoring this cross-sectional dependence within units by incorporating it into the error term will lead to an inconsistent estimate (Hoyos & Sarafidis, 2006). To check the cross-sectional dependency within individuals (countries), Pesaran CD test and Breusch-Pagan test (Breusch & Pagan, 1980) are mostly used (Bhattacharya et al., 2016). In this study, the number of individuals (13) is less than the time dimension (26 years). In that case, instead of Pesaran CD test, Breusch-Pagan test used to check for cross-sectional dependence within countries is a better fit. If the hypothesis of the cross-dependency within variables is verified, the panel unit root test will be applied using the second-generation test suggested by Pesaran (2007). The null hypothesis is that there is homogenous non-stationarity. If one fails to reject the null hypothesis of the existence of panel unit root, it means the variances of the time series are unstable over time leading probably to the long-run relationships between variables. In that case, the panel error correction model (PVCMM) will be useful. The PVCMM equations are stated as follows:

$$\begin{bmatrix} \Delta \ln AgVA_{it} \\ \Delta \ln CO_{2it} \\ \Delta \ln REC_{it} \\ \Delta Trade_{it} \\ \Delta GFCF_{it} \end{bmatrix} = \begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \alpha_3 \\ \alpha_4 \\ \alpha_5 \end{bmatrix} + \sum_{k=1}^p \begin{bmatrix} \mathcal{G}_{11p} & \mathcal{G}_{12p} & \mathcal{G}_{13p} & \mathcal{G}_{14p} & \mathcal{G}_{15p} \\ \mathcal{G}_{21p} & \mathcal{G}_{22p} & \mathcal{G}_{23p} & \mathcal{G}_{24p} & \mathcal{G}_{25p} \\ \mathcal{G}_{31p} & \mathcal{G}_{32p} & \mathcal{G}_{33p} & \mathcal{G}_{34p} & \mathcal{G}_{35p} \\ \mathcal{G}_{41p} & \mathcal{G}_{42p} & \mathcal{G}_{43p} & \mathcal{G}_{44p} & \mathcal{G}_{45p} \\ \mathcal{G}_{51p} & \mathcal{G}_{52p} & \mathcal{G}_{53p} & \mathcal{G}_{54p} & \mathcal{G}_{55p} \end{bmatrix} \times \begin{bmatrix} \Delta \ln AgVA_{it-1} \\ \Delta \ln CO_{2it-1} \\ \Delta \ln REC_{it-1} \\ \Delta Trade_{it-1} \\ \Delta GFCF_{it-1} \end{bmatrix} + \begin{bmatrix} \delta_1 \\ \delta_2 \\ \delta_3 \\ \delta_4 \\ \delta_5 \end{bmatrix} ECT_{it-1} + \begin{bmatrix} \mu_{1it} \\ \mu_{2it} \\ \mu_{3it} \\ \mu_{4it} \\ \mu_{5it} \end{bmatrix} \quad (3)$$

In equation (3), Δ stands for the first difference of each variable. The long-term relationship between variables is captured by the error correction term and represented by ECT_{it-1} . In equation (3), μ stands for the random error term.

The study covers 13 countries in West Africa, and the data from the World Bank Indicator (WDI, 2020) are used. The economic growth was, on average, about 3.86% per year. The agricultural sector was the main driver of economic growth and reached about 61.41% of GDP. It accounted on average for 30.38% of gross domestic product (GDP) throughout 1990–2015 (Table 1).

The renewable energy consumption (REC) in the region has reached an average of 72.17% of total final energy consumed in West African countries. This would significantly influence agricultural development through direct use for agricultural production and reduce CO₂ emissions (climate change). CO₂ emissions, which are the main components of greenhouse gas, could also directly or indirectly affect the region's economic growth. The average CO₂ emissions in the selected countries reached a maximum of 106 thousand kilotons. The trade openness reached about 62.07% of GDP.

Nigeria and Ghana, the leading economies of the region, have the highest economic growth (5.11% and 5.38%, respectively), while Guinea Bissau, Ivory Coast, and Sierra Leone have the lowest rate of economic growth (2.33%, 2.55%, and 2.82%, respectively). The average agricultural value-added in the leading economy of West Africa reached annually 24.81% for Nigeria and 33.22% for Ghana. The agricultural sector has an important contribution to GDP in the overall selected countries with more significance in Sierra Leone (49.02% of GDP) followed by Guinea-Bissau (47.39% of

Table 1. Average annual of the variables used in the study (1990–2015)

Countries	GDP	AgVA	REC	Trade	GFCF	CO ₂
Benin	4.52	27.18	69.09	55.94	20.50	2852.46
Burkina Faso	5.36	32.52	85.25	39.31	21.87	1344.14
Ivory Coast	2.55	25.12	72.49	78.11	11.75	6963.26
Gambia	3.33	21.58	58.13	64.51	15.10	295.75
Ghana	5.38	33.22	64.47	75.98	20.79	7728.04
Guinea	3.75	19.04	82.92	61.40	20.50	1795.66
Guinea-Bissau	2.33	47.39	88.54	49.55	15.89	203.82
Mali	4.43	34.79	80.36	55.59	19.17	794.16
Mauritania	3.83	28.29	40.16	92.46	30.50	1581.00
Nigeria	5.11	24.81	86.29	38.62	29.62	73652.05
Senegal	3.58	15.64	47.89	61.04	21.33	5055.58
Sierra Leone	2.82	49.02	85.99	53.41	11.71	640.45
Togo	3.22	35.36	76.68	80.91	16.19	1532.21
West Africa	3.86	30.38	72.17	62.07	19.61	8033.74

GDP). However, Burkina Faso has the highest agricultural value addition (32.52%) and economic growth within the West African Economic Monetary Union (WAEMU), a sub-regional group of West Africa, sharing the common currency (the Franc CFA). The proportion of REC of the total energy used in West Africa was relatively high in countries such as Sierra Leone (85.99%), Nigeria (86.29%), Guinea-Bissau (88.54%), Mali (80.36%), and Burkina Faso (85.25%). The gross fixed capital formation (GFCF) that captures the investment level reached on average 30.50% of GDP in Mauritania, while it was only about 11.71% in Sierra Leone over the study period. However, Nigeria has the highest level of CO₂ emissions (73652.05 kt) and low level observed in Guinea-Bissau (203.82 kt). This can be explained by the level of economic development of each country, as highlighted by Kuznets (1955).

3. RESULTS

The correlation between variables and the multicollinearity test are provided in Table 2. The average variance inflation factors ($VIF = 1.31$) are less than 5% showing that multicollinearity between the considered variables is not a problem in the estimation process (O'Brien, 2007). Table 2 shows a positive correlation between REC and agricultural development (AgVA). The econometrics tests would shed light on the causality between these variables. However, a high and negative correlation between REC and trade openness was observed (Table 2).

Table 2. Correlation between variables and multicollinearity test

Variables	AgVA	REC	CO ₂	GFCF	Trade
AgVA	1	–	–	–	–
REC	0.49	1	–	–	–
CO ₂	–0.18	0.16	1	–	–
GFCF	–0.24	–0.26	0.18	1	–
Trade	–0.17	–0.52	–0.26	0.20	1
VIF (in %)	–	1.45	1.16	1.17	1.48
Mean VIF	1.31				

Moreover, there was a negative and low correlation between gross fixed capital formation and REC. This suggests a probable substitution between the use of renewable energy and investment captured by gross fixed capital formation. Given the individuals and the time dimension, the Breusch-Pagan test was a better fit to check the existence of cross-sectional dependence within countries (Table 3).

Table 3. Breusch-Pagan test of independence

Variables	Chi-squared statistics	p-value
AgVA	330.124***	0.0000
REC	768.506***	0.0000
CO ₂	1422.923***	0.0000
GFCF	299.576***	0.0000
Trade	242.946***	0.0000

Note: (***) indicates the significance level at 1%.

The results show that the probability values of all variables are less than 1% level. This implies that the hypothesis of cross-sectional dependency

within individuals cannot be rejected. Therefore, the first generation test of panel unit roots is no longer appropriate. Thus, the second-generation panel unit root tests will be applied. The author showed a cross-sectional dependence between individuals. In that case, the second-generation test of panel unit root proposed by Pesaran (2007) is the most appropriate (Table 4).

Table 4. Pesaran panel unit root test with cross-sectional

Variable	CIPS	Critical value		
		10%	5%	1%
lnAgVA	-2.744	-2.66	-2.76	-2.96
lnREC	-2.173	-2.14	-2.25	-2.45
lnCO ₂	-2.355	-2.14	-2.25	-2.45
GFCF	-2.805	-2.66	-2.76	-2.96
lnTrade	-2.132	-2.14	-2.25	-2.45

The null hypothesis states that all variables are homogeneous non-stationary. The alternative hypothesis is that the time series is stationary, and the integration of variables is no longer important. The results show that the Pesaran statistic values (CIPS) are all greater than the critical values for all variables at least at 1% level. This result suggests that all variables need to be integrated because they are non-stationary at the level. To check the existence of the long-run relationship between variables, Westerlund panel cointegration tests were used (Westerlund, 2007). This test is more appropriate than the cointegration test performed by Perdoni (1999) since there is a cross-sectional dependence between individuals. This test assumes that there is no cointegration between variables. The results show that all statistics are significant, at least at a 5% level, including the robust *p*-values (Table 5).

Table 5. Westerlund ECM panel cointegration tests

Variable	Statistics	Value	Z-value	Robust <i>p</i> -value
AgVA	Gt	-5.008***	-11.908	0.000
	Ga	-31.848***	-10.814	0.000
	Pt	-18.209***	-12.336	0.000
	Pa	-31.641***	-13.695	0.000
REC	Gt	-3.861***	6.756	0.000
	Ga	-27.839***	8.641	0.010
	Pt	-13.099**	6.384	0.020
	Pa	-22.964**	8.457	0.040

Note: (***) indicates the significance level at 1%, (**) is the significance level at 5%.

This implies that the null hypothesis of the absence of cointegration can be rejected, hence the existence of a long-run cointegration among variables. However, the stability of this long-run relationship should be tested. The Granger causality test results show a unidirectional causality, running from REC to agricultural value-added (Table 6).

The Granger causality test shows that a bidirectional hypothesis is verified between agricultural value-added and trade openness. Therefore, any change in trade openness will affect agricultural value-added and vice-versa. On the one hand, this result can be explained by the fact that most agricultural products in West Africa are exported as raw materials (cotton, cocoa, coffee, cashew, rubber) and the subject of important revenues. Alternatively, West African countries are net importers of most transformed agricultural goods, and trade openness can facilitate transactions. This result was also supported by Ben Jebli and Ben Youssef (2015). Similar results were found in the literature (Raeni et al., 2019; Brini et al., 2017; Marques & Fuinhas, 2011) but contradict Liu et al. (2017) who show a positive relationship between REC and trade openness.

Raeni et al. (2019) found no causality between trade and REC in the Iranian case study, while Marques and Fuinhas (2017) found that the market was not a determinant of renewable energy use in the case study of 24 European countries. The results also showed no causality between trade openness and CO₂ emissions. This result is similar to Raeni et al. (2019), but contradicts Ben Jebli and Ben Youssef (2015) who found that trade openness may be a source of global warming as a transaction in transportation traded goods may lead to more CO₂ emissions. The results show that the long-run coefficients of agricultural value-added, REC, and trade openness are significant at a 1% level (Table 7). It implies that the three models out of five have a long-run relationship.

The negative sign is associated with agricultural value-added and trade openness, while it is positive for the REC. This suggests that trade openness will not favor agricultural development in West African countries in the long run. The adjustment amount of REC from short-run to long-run is 0.076. This suggests that the previous shocks on agricultural value-added, CO₂ emissions, gross

Table 6. Granger causality test

Variable	AgVA	REC	CO ₂	Trade openness	GFCF
AgVA	–	Unidirectional causality between <i>REC</i> and <i>AgVA</i>	Unidirectional causality between <i>CO₂</i> and <i>AgVA</i>	Bidirectional causality between <i>Trade openness</i> and <i>AgVA</i>	No causality
REC	Unidirectional causality between <i>REC</i> and <i>AgVA</i>	–	No causality	No causality	Unidirectional causality between <i>GFCF</i> and <i>REC</i>
CO ₂	Unidirectional causality between <i>CO₂</i> and <i>AgVA</i>	No causality	–	No causality	No causality
<i>Trade openness</i>	Bidirectional causality between <i>Trade openness</i> and <i>AgVA</i>	No causality	No causality	–	Unidirectional causality between <i>GFCF</i> and <i>Trade openness</i>
GFCF	No causality	Unidirectional causality between <i>GFCF</i> and <i>REC</i>	No causality	Unidirectional causality between <i>GFCF</i> and <i>Trade openness</i>	–

fixed capital formation, and trade openness will increase the demand for renewable energy by 7.6% in the long run. This can be achieved by increasing the investment in renewable energy production significantly (Bayale et al., 2021; Le & Van, 2020; Liu et al., 2017; Rafindadi & Ozturk, 2017). The evidence is that there is a unidirectional relationship between GFCF and REC in the short run, indicating that the GFCF is a determinant of renewable energy use. However, the results indicate that any shocks in the previous periods could have negatively resulted in agricultural growth and

trade openness at the long-run equilibrium. These shocks might be climate change (increase in CO₂ emissions), low investment in renewable energy production, or the pollution from globalization that is prone to trade openness (Schwerhoff & Sy, 2017; Ben Jebli & Ben Youssef, 2017).

4. DISCUSSION

This result suggests that any REC change has an immediate impact on agricultural development

Table 7. Panel Vector Error Correction Model (PVCM) estimating long-run causality

Variables	$\Delta(\ln AgVA)$	$\Delta(\ln CO_2)$	$\Delta(\ln REC)$	$\Delta(\ln Trade)$	$\ln GFCF$
Long-run					
EC_{t-1}	-1.099*** (0.082)	0.161 (0.123)	0.076*** (0.036)	-0.506*** (0.146)	3.996 (3.943)
Short-run					
$\Delta(\ln AgVA_{t-1})$	0.053 (0.056)	-0.100 (0.084)	-0.033322 (0.025)	0.340*** (0.100)	-1.572 (2.702)
$\Delta(\ln CO_{2,t-1})$	-0.100*** (0.040)	-0.470*** (0.060)	-0.019 (0.018)	-0.0096 (0.071)	-1.274 (1.932)
$\Delta(\ln REC_{t-1})$	-0.467*** (0.121)	0.037 (0.182)	-0.518*** (0.054)	-0.313 (0.216)	-4.496 (5.808)
$\Delta(\ln Trade_{t-1})$	0.163*** (0.030)	-0.044 (0.046)	-0.015 (0.013)	-0.444*** (0.054)	-0.155 (1.467)
$\ln GFCF_{t-1}$	0.001 (0.001)	-0.002 (0.001)	0.001*** (0.001)	-0.004** (0.002)	-0.192*** (0.058)
C	0.002 (0.006)	-0.010 (0.009)	-0.001 (0.002)	0.0001 (0.011)	0.146 (0.307)
R ²	0.5119	0.2083	0.2825	0.3579	0.4550
Adj. R ²	0.5019	0.1920	0.2678	0.3447	0.2590

Note: (***) indicates the significance level at 1%, (**) is the significance level at 5%.

in the West African region. Similar results were found by Raeeni et al. (2019) in the Iranian case study and by Ouedraogo (2013) in WAEMU countries. In achieving sustainable development goals in West Africa, the role of RE in agricultural development recalls attention on financing RE (Raeeni et al., 2019; Schwerhoff & Sy, 2017). Moreover, the results indicate that the unidirectional hypothesis is verified between AgVA and CO₂ emissions. This causality runs from CO₂ emissions to agricultural value-added. This result implies that climate change induced by an increase in CO₂ emissions could directly impact agricultural value-added. This result corroborates most studies recognizing the pronounced impacts of climate change on crop productivity in West African countries (Ali et al., 2020; Parkes et al., 2018; Ali, 2018). However, the bidirectional hypothesis was verified in the case study by Ben Jebli and Ben Youssef (2015) in Tunisia. The results show that the unidirectional hypothesis is verified between REC and gross fixed capital formation (GFCF) running from GFCF to REC, on the one hand, and trade openness and GFCF running from GFCF to trade openness, on the other hand. This implies that any change in investment level will affect renewable energy consumption and trade openness. Indeed, an increase in investment level might increase renewable energy technologies adoption, therefore, requiring trade openness. This result is similar to those from Le and Van (2020).

In light of the above results, there is a need to test the stability of the long-run relationship between variables, as shown by the inverse roots of AR characteristic polynomial (Figure 1).

Figure 1 shows that all inverse roots lie within the unit circle. It implies that the long-run relationship between variables is stable; therefore, any energy policy response in this study could be validated.

The results from PVCM indicate that CO₂ emissions, renewable energy consumption, and trade openness at lag one significantly influence agricultural value-added. Renewable energy consumption might favor agricultural growth, while CO₂ emissions and trade openness would negatively affect. Trade openness in the context of globalization could facilitate technology transferability and input supply for agricultural development. It could also be a catalyst to the agricultural commodity market at the international level. Indeed, increasing the demand for renewable energy can increase agricultural production, and trade openness can be a catalyst in trading agricultural commodities and technology transfer for agricultural development. The negative sign might be explained by the fact that most of the agricultural products subject to international trade are exported as raw materials. The transformed goods from these raw materials are then imported for households' final consumption. In that case, importing more agricul-

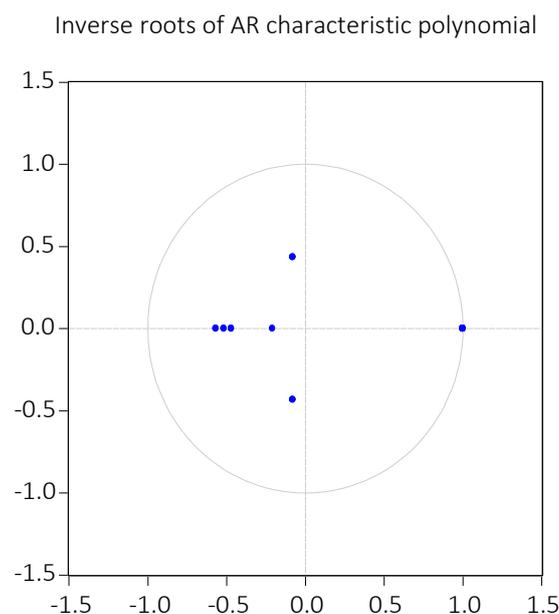


Figure 1. Test of stability of long-run relationship among variables

tural commodities for households' consumption will favor mostly the agricultural development of net exporting countries while negatively affect the agriculture growth of net importer countries with low financial development (Kim et al., 2012). The other reason is sanitary and phytosanitary barriers that mostly face the agricultural commodities trade. Also, agricultural production subsidies in developed countries can lead to the negative impact of trade openness on agricultural development in developing countries, including the West

African countries. Reducing trade barriers can result in the expected impact of regional integration on countries' economies.

The second and fifth models indicate that there is no long-run causality relationship. The results show that agricultural value-added, renewable energy consumption, trade openness, and gross fixed capital formation do not significantly affect CO₂ emissions. Similarly, the agricultural value-added, renewable energy consumption, trade

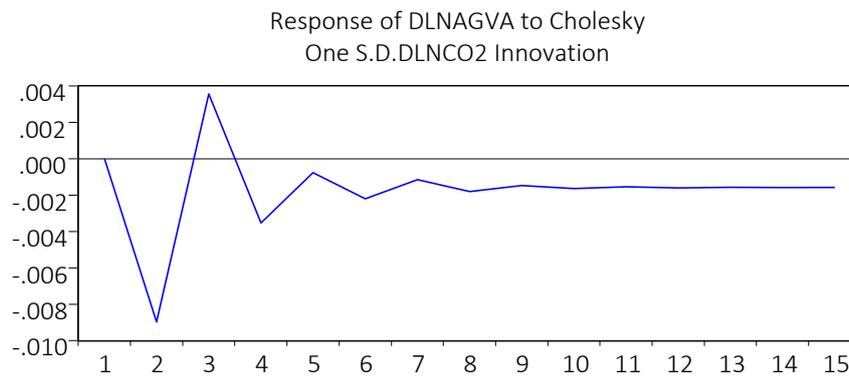


Figure 2a. CO₂ policy response on agricultural development by 2030 (SDG agenda)

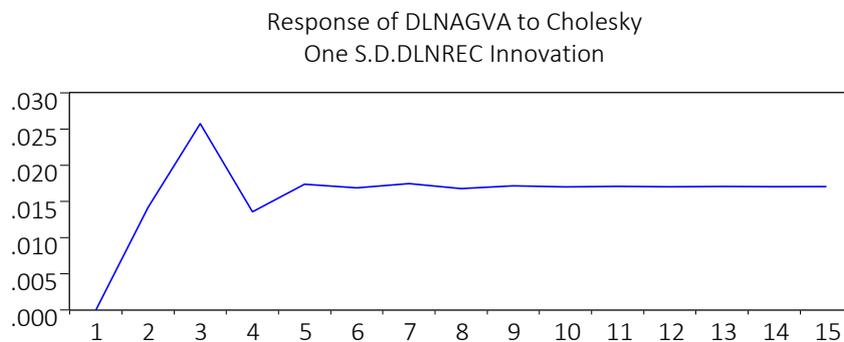


Figure 2b. REC policy response on agricultural development by 2030 (SDG agenda)

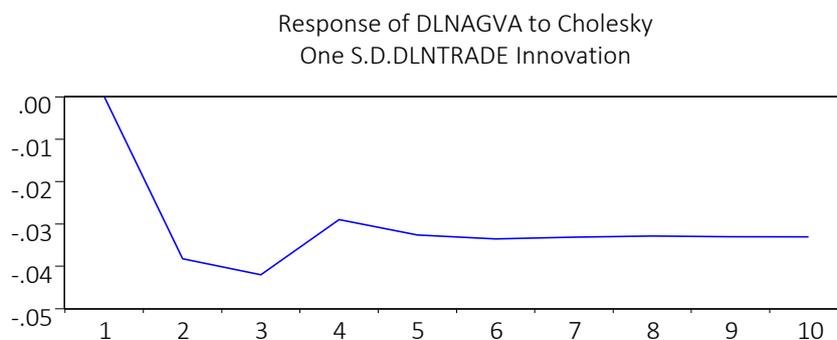


Figure 2c. Regional integration policy response on agricultural development by 2030 (SDG agenda) in West Africa

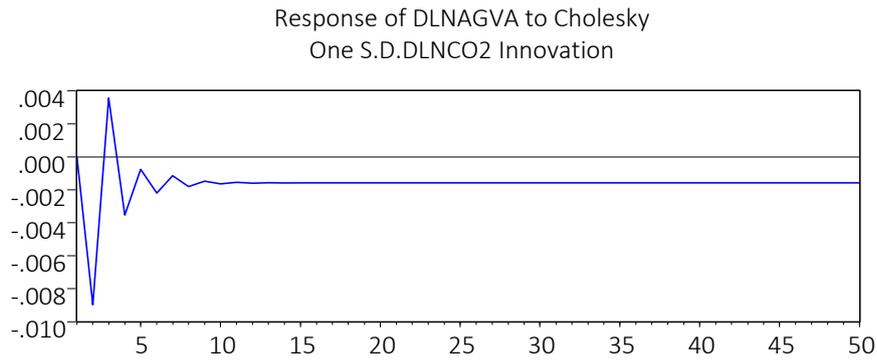


Figure 3a. CO₂ emissions policy response on agricultural development by 2063 (African Union agenda)

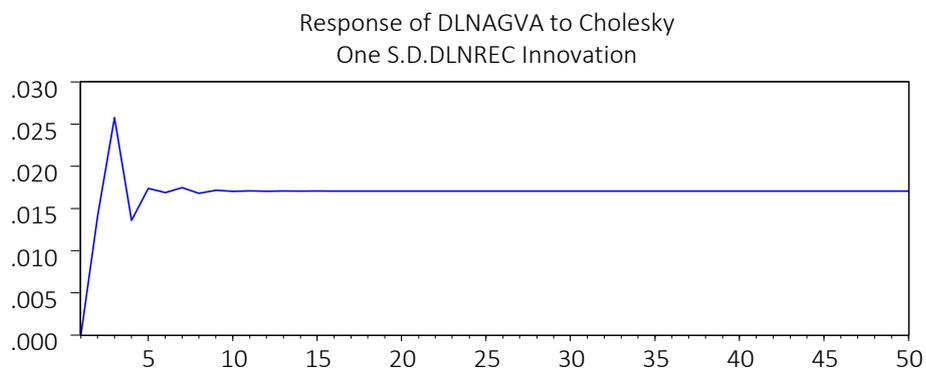


Figure 3b. REC policy response on agricultural development by 2063 (African Union agenda)

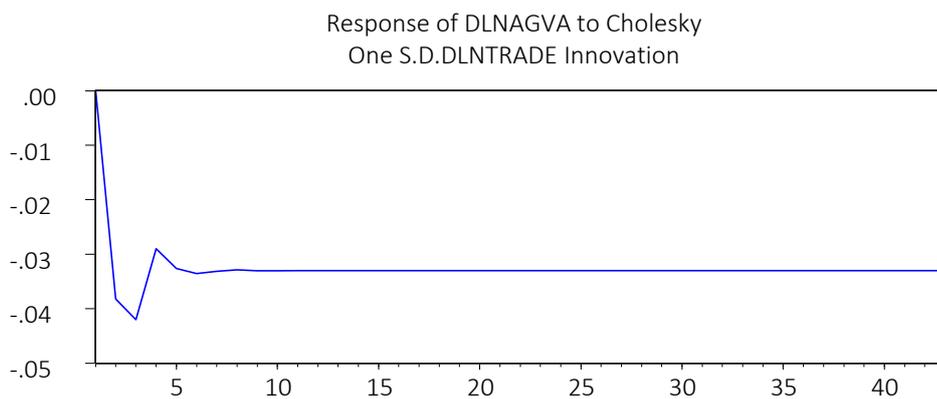


Figure 3c. Trade openness policy response on agricultural development by 2063 (African Union agenda)

openness, and CO₂ emissions do not significantly affect gross fixed capital formation. However, the findings show that agricultural growth has positively affected trade openness, while there was a substitution between gross capital formation and trade openness.

In the specific context of West African countries, it becomes important to analyze the policy response of Sustainable Development Goals agenda (objective number 7) that seeks to ensure access to affordable, reliable, sustainable, and modern energy for all by 2030 with a focus on clean, renewable

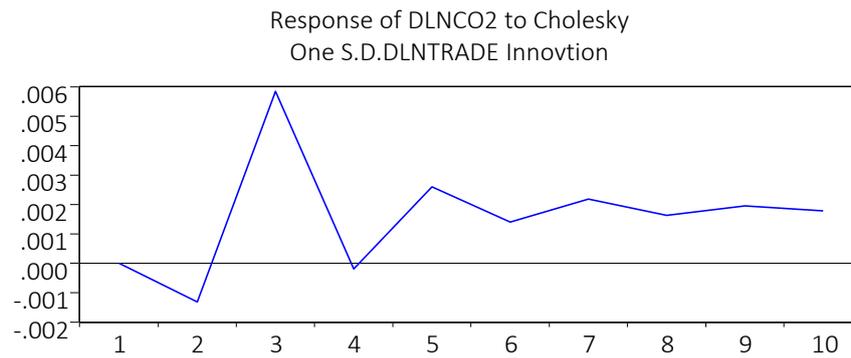


Figure 4a. Trade openness policy response on CO₂ emissions by 2030 (SDG agenda)

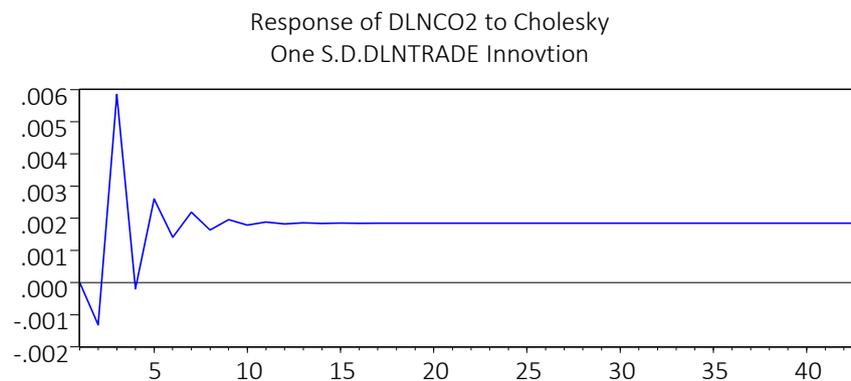


Figure 4b. Trade openness policy response on CO₂ emissions by 2063 (African Union agenda)

energy use in combating climate change. Similarly, Agenda 2063 of the African Union has set a roadmap of environmentally sustainable and climate-resilient economies by increasing the use of renewable energy and fostering regional integration. Either SDG agenda (Figures 2a, 2b, and 2c) or Agenda 2063 of the African Union (Figures 3a, 3b, and 3c), the impulse responses of an increase in CO₂, REC and trade openness have similar trends. Figures 2a, 2c, 3a, and 3c show that an increase in CO₂ and trade openness considering the SDG agenda and African Union agenda will negatively affect the agricultural value-added of West African countries starting from 2023. However, an increase in REC will positively affect agricultural development. The high effect of the use of REC will be observed for the first three years and will start declining in the fourth year. After 5 years, the impact of REC on agriculture will remain constant but positive.

The results show that globalization through trade openness will increase CO₂ emissions. Whether one considers the SDG agenda (Figure 4a) or the

African Union agenda (Figure 4b), similar trade openness effects on CO₂ emissions are observed. These results corroborate those from Mutascu and Sokic (2020) and Sannasse and Seetanah (2016). The positive relationship between trade openness and pollution was also found by Bataka (2021) who found that globalization contributes to environmental pollution in Sub-Saharan African countries.

Trade openness can increase CO₂ emissions if most imported goods are highly pollutants. Also, it is well recognized that developing countries like the West African countries have dirty industries with non-binding environmental laws in contrast to developed countries. These findings suggest that the impact of trade openness on CO₂ emissions may depend on economic development; therefore, the re-examination of trade policy in developing countries toward a clean environment is needed for sustainable economic development. Findings also show that trade openness can increase renewable energy consumption, which can be used for agricultural production (see Figures A1 and A2 in Appendix A).

CONCLUSION

The use of renewable energy is a key strategy in combating climate change, one of the most world's concerns of the century. In the context of developing countries where rapid population growth and urbanization are expected, renewable energy consumption (REC) must be encouraged to meet the increasing demand for energy while reducing CO₂ emissions. Therefore, the new empirical evidence is needed for setting a roadmap of regional energy policy for economic development. However, the specific energy policy could depend not only on the type of hypothesis of the relationship between REC and economic development but also on the geographical position of the region and the main drive of countries' economies. In the context of developing countries, including the West African region, where the agricultural sector remains the main driver of countries' economies, not much attention has been given to the role of trade openness as a catalyst for RE technology transfer and REC in investigating the dynamic between CO₂ emissions as the main source of climate change and agricultural development of the region. This study analyzes the relationship between CO₂ emissions and agricultural development in West African countries by focusing on the role of renewable energy and trade openness.

The second-generation panel unit root tests, the Westerlund cointegration methods were used with the data from 13 countries of West Africa from 1990 to 2015. A panel error correction model was used to analyze the long-run relationship between variables. A panel Granger causality test was also used to check the causality direction between variables. Findings show a unidirectional relationship between agricultural value-added (AgVA) and CO₂ emissions running from CO₂ emissions to AgVA. The unidirectional causality was also found from REC to agricultural value-added. The results confirm the unidirectional hypothesis running from the gross fixed capital formation (GFCF) to REC, on the one hand, and from GFCF to trade openness, on the other hand. Moreover, the bidirectional hypothesis is verified between agricultural development and trade openness with positive and significant effects. The results show that previous shocks on different variables might result in a negative effect on agricultural value-added and trade openness in the long-run. However, the results show that the previous shocks on agricultural value-added, CO₂ emissions, gross fixed capital formation, and trade openness will increase the renewable energy demand by 7.6% in the long run. Considering the SDG agenda (Agenda 2030) or the African Union agenda (Agenda 2063), the impulse response of REC showed a positive effect on agricultural value-added while negatively related to CO₂ emissions and trade openness. Increasing the demand for renewable energy can spur agricultural production, and trade openness can ease the trade of agricultural commodities. However, exporting more agricultural commodities as raw materials and importing mostly high pollutant commodities will result in a positive effect relationship between trade openness and CO₂ emissions, as shown in the policy responses of Agenda 2030 and 2063 in the results. These results suggest that fostering renewable energy policy in West African countries will contribute to agricultural development. However, a re-examination of trade policy to reduce environmental pollution should be a priority for the West African countries to gain from the regional integration.

AUTHOR CONTRIBUTIONS

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Formal analysis: Essossinam Ali.

Funding acquisition: Essossinam Ali.

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Methodology: Essossinam Ali.

Project administration: Essossinam Ali.

Resources: Essossinam Ali.

Software: Essossinam Ali.

Supervision: Essossinam Ali.

Validation: Essossinam Ali.

Visualization: Essossinam Ali.

Writing – original draft: Essossinam Ali.

Writing – review & editing: Essossinam Ali.

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APPENDIX A

Response to Cholesky One S.D. Innovations

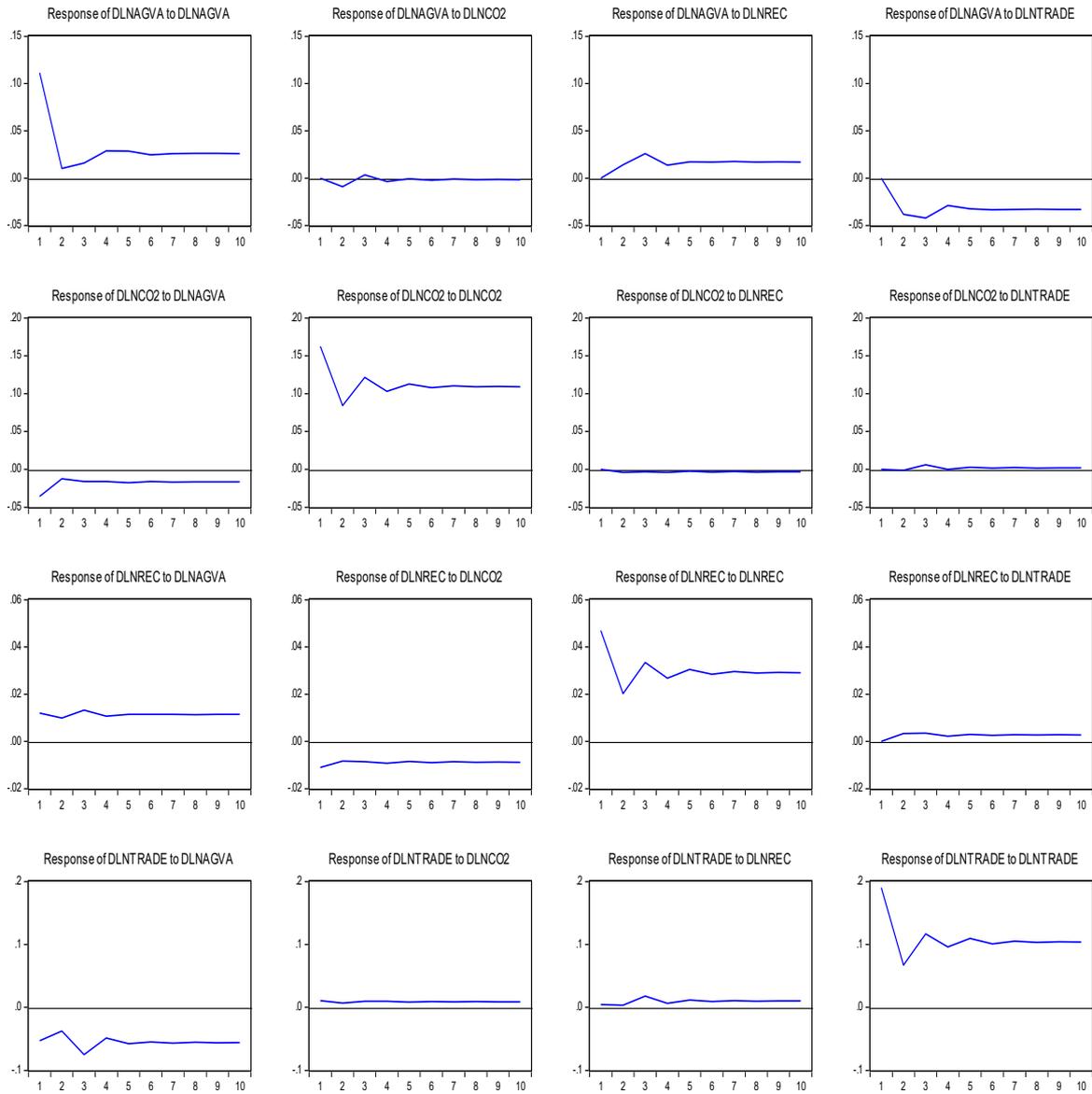


Figure A1. Policy responses in light of sustainable development agenda (Agenda 2030)

Response to Cholesky One S.D. Innovations

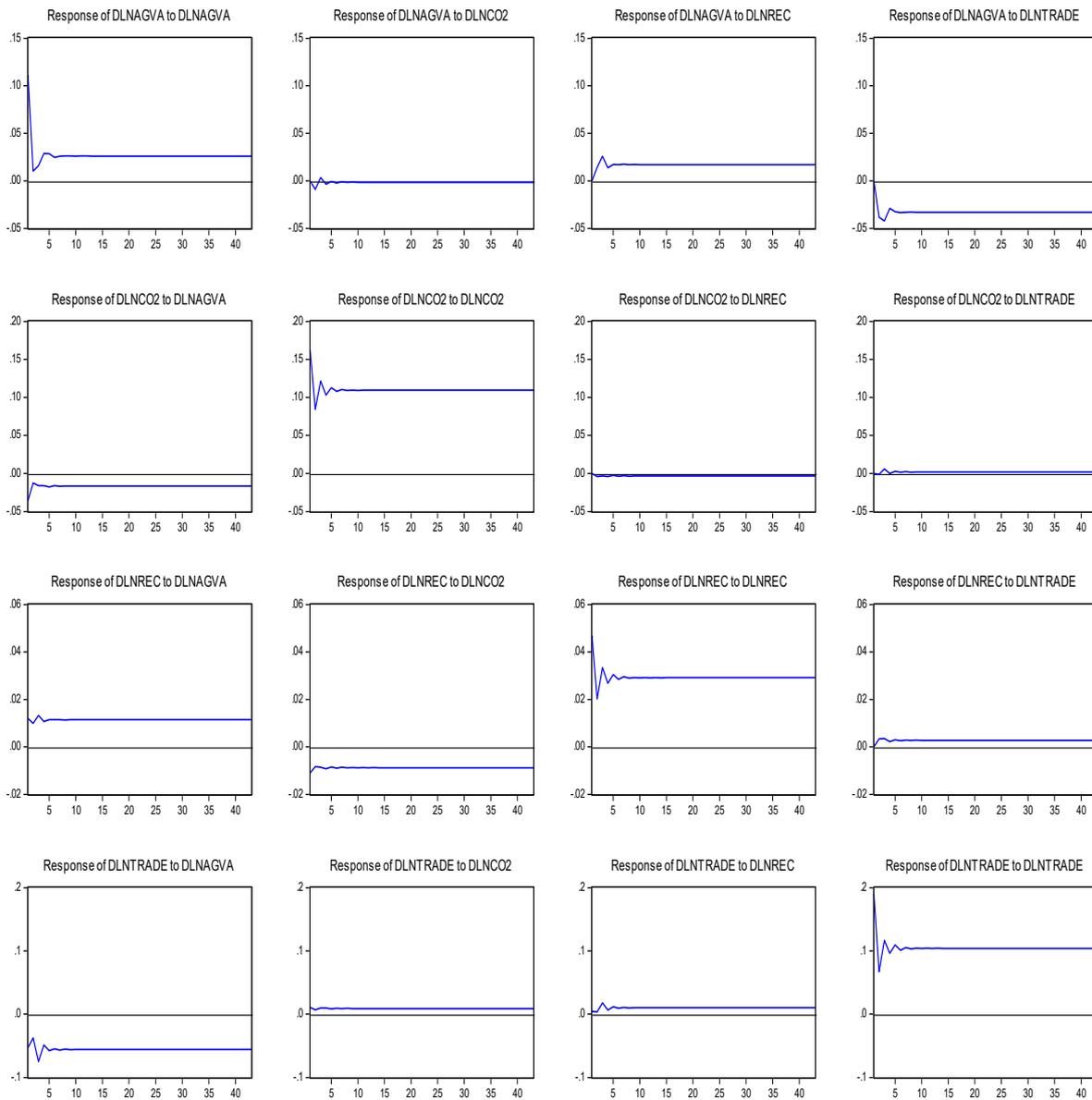


Figure A2. Policy response in light of African Union agenda (Agenda 2063)